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#### **ORIGINAL ARTICLE**

## Low molecular weight heparin (LMWH) for primary thrombo-prophylaxis in patients with solid malignancies – systematic review and meta-analysis

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#### ABSTRACT

**Background.** Patients receiving chemotherapy for cancer are at increased risk for venous thromboembolism (VTE). We performed a meta-analysis of all randomized controlled trials (RCTs) which evaluated low molecular weight heparin (LMWH) as primary prophylaxis in ambulatory patients with solid malignancies.

**Methods.** A comprehensive search was conducted until October 2013. Primary outcome was symptomatic VTE. Secondary outcomes were pulmonary embolism (PE), any VTE, deep vein thrombosis (DVT), mortality and adverse events.

**Results.** Eleven trials met the inclusion criteria, and evaluated a total of 6942 patients. Primary prophylaxis with LMWH reduced symptomatic VTE (RR 0.46, 95% CI 0.32–0.67) and the rate of PE (RR 0.49, 95% CI 0.29–0.84). In the subgroup analysis of VTE in patients with lung and pancreatic cancers LMWH further reduced VTE [RR 0.42 (95% CI 0.25–0.71); RR 0.31 (95% CI 0.18–0.55), respectively]. Meta-analysis of six trials which reported survival outcomes revealed no statistically significant benefit for LMWH in one-year mortality rates (RR 0.93, 95% CI 0.83–1.04). There was no significant increase in major bleeding events (RR 1.28, 95% CI 0.84–1.95).

**Conclusions.** LMWH reduces the incidence of symptomatic VTE and PE in patients receiving chemotherapy for cancer, with no apparent increase in major bleeding. The benefit is most apparent in pancreatic cancer and also lung cancer. VTE prophylaxis should be considered for these specific populations.

Venous thromboembolism (VTE) is one of the principal causes of morbidity and mortality in cancer patients. It occurs in 4–20% of cancer patients, and it is one of the leading causes of deaths [1,2]. The risk of venous thromboembolic events, including pulmonary embolism (PE) and deep vein thrombosis (DVT) in cancer patients varies according to disease-specific factors, such as site, stage, and type of malignancy [3,4]. The incidence of VTE is higher in pancreas, stomach and lung cancers [5]. VTE risk is further increased by anti-cancer therapies, with significant increases in VTE in cancer patients treated with chemotherapy and hormonal-therapy [6–8]. Several studies have implied that low molecular heparin (LMWH) may confer anti-neoplastic

properties, such as inhibiting pathways of angiogenesis, cellular adhesion, and tumor invasion and by inducing apoptosis and hence may affect survival not only by reducing VTE [9-12].

Several studies have evaluated the role of LMWH as primary thromboprophylaxis in cancer patients, yielding inconsistent results. The general recommendation established by four organizations [American College of Chest Physicians (ACCP), American Society of Clinical Oncology (ASCO), National Comprehensive Cancer Network (NCCN) and European International Good Clinical Practices Guidelines (GCPG)] is that routine pharmacologic thromboprophylaxis is not currently recommended in cancer outpatients, based upon moderate evidence. Nevertheless, recent ASCO

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guidelines recommend considering primary LMWH prophylaxis based on a case-by-case basis in highly selected outpatients with solid tumors receiving chemo-therapy [5,13,14].

In view of the conflicting data from randomized clinical trials, and due to the different biological course and mechanism of certain malignancies, which may result in higher efficacy for LMWH in some subpopulations, we performed a systematic review of the literature and a meta-analysis of all randomized trials to evaluate the impact of LMWH primary prophylaxis onVTE incidence as well as survival in cancer patients. Moreover, we aimed to assess subpopulations of specific malignancies considered as high risk for developing VTE.

### Methods

#### Data sources

We searched the Cochrane Central Register of Controlled Trials, published in The Cochrane Library, PubMed (1966-Oct 2013); the database of clinical trials in cancer patients; conference proceedings of the American Society of Clinical Oncology (1995-Oct 2013), American Society of Hematology (2006-Oct 2013), proceedings of the European Society of Medical Oncology (ESMO) (2006-Oct 2013) and the European Hematology Association (EHA) (2006-Oct 2013); and databases of ongoing and unpublished trials: http://www.clinicaltrials.gov and http://www.clinicaltrials.nci.nih.gov. The terms (tumor OR malign\* OR carcinoma\* OR cancer) AND (heparin OR low-molecular weight heparin OR enoxaparin OR dalteparin OR reviparin OR certoparin OR tinzaparin OR bemiparin OR nadroparin OR \*parin) AND (thromboembolism) were crosssearched. The result was limited to randomized controlled trials using a highly sensitive filter [15]. We scanned references of all included trials and reviews identified for additional studies.

#### Study selection

We included randomized controlled trials that compared the addition of LMWH to standard chemotherapy in ambulatory cancer patients, as primary thromboprophylaxis. We included trials regardless of publication status, date of publication, and language. Two authors (IBA and AG) independently inspected each reference title identified by the search and applied the inclusion criteria.

#### Data extraction and quality assessment

Trials that fulfilled the inclusion criteria were assessed for methodological quality by two authors (IBA and AG). Both reviewers independently assessed risk of bias in the included trials. We used the Cochrane Collaboration's tool for assessing risk of bias. We individually assessed the following domains: random sequence generation, allocation concealment, blinding of participants and personnel, blinding of outcome assessment, incomplete outcome data reporting and selective outcome reporting. We separately assessed each domain and graded it as low risk for bias, unclear risk (lack of information or uncertainty over the potential for bias), or high risk for bias according to the criteria specified in the Cochrane Handbook version 5.1.0 [15]. The same two authors independently extracted the data from publications of included trials. The data extraction was discussed, decisions were documented, and, if necessary, the authors of the trials were contacted for clarification. Authors of included trials were contacted for all data relevant to the primary and secondary outcomes (VTE, survival and safety data) of the study and quality variables. In case of several publications for the same trial, the most updated one was extracted. Safety outcomes were pooled from the most updated publication of every trial.

#### Outcome measures

The primary outcome was symptomatic VTE, which was defined according to a unified definition in all included trials. Secondary outcomes were any VTE, PE, DVT, all-cause mortality, and toxicity (defined as grade 3 or 4 hematological and non-hematological adverse events). Regarding bleeding, we extracted data regarding major bleeding and clinically relevant bleeding (defined as major plus minor bleeding).

#### Data synthesis and statistical analysis

Risk ratios (RRs) and 95% confidence intervals (CIs) for dichotomous data were estimated using the Mantel-Haenszel method and pooled according to inverse of variance method [Review Manager (RevMan), version 5.1 (Copenhagen: The Nordic Cochrane Centre, The Cochrane Collaboration, 2011)]. A RR less than 1.0 was in favor of LMWH therapy. We assessed heterogeneity of trial results by calculating a  $\chi^2$ -test of heterogeneity and the I<sup>2</sup> measure of inconsistency. We chose a random-effects model and used the Der Simonian and Laird method for all analyses due to different types of cancer and therefore different VTE risks [16]. Subgroup analyses were performed for lung cancer and pancreatic cancer, which are regarded as cancers with a high thrombogenic potential.

For calculating number needed to treat (NNT) in order to evaluate the additive effect of LMWH on

the absolute risk for VTE, we retrieved the data on in each arm (LMWH vs. control). Risk was calculated by multiplying the absolute risk of the control arm by (1-RR for VTE/PE/DVT).

#### Results

The literature search identified 747 trials up to October 2013, of which 69 were considered potentially relevant. Additional trials were identified by searching conference proceedings and electronic resources of ongoing trials. Figure 1 illustrates the process of study selection. Eleven trials were designed to evaluate the effect of primary LMWH prophylaxis on cancer-related VTE and fulfilled the inclusion criteria for published studies [17–27; including safety reports]. The TOPIC trial was analyzed separately for breast cancer patients and for lung cancer patients [26,27].

#### Studies characteristics

A total of 6942 patients were randomly assigned in the 11 trials included in the meta-analysis for VTE. Not all 11 reported the same outcomes, hence analysis has been performed upon specific outcome. One trial was in the form of an abstract (CONKO004, 22). Six studies [17–20,23,24] were included in the all-cause mortality analysis (12 months) and evaluated 2550 patients.

#### Trial design

In three trials [19,21,23] patients were randomly assigned to nadroparin in addition to standard-ofcare therapy. Four trials used dalteparin [17,18,20,25], two trials evaluated certoparin [26,27], one trial used



Figure 1. Randomized controlled trials search and selection.

enoxaparin [22] and one semuloparin [24]; patient population varied as reflected in Table I.

#### Quality of trials

Allocation concealment was reported as adequate in seven trials [18,19,21,23,25,26] and was not reported in the other four trials. Few of the trials have been open-labeled. The quality assessment of the included trials is described in detail in Supplementary Table I (available online at http://informahealthcare. com/doi/abs/10.3109/0284186X.2014.934397). We appraised the rate of patients loss to follow-up and in the majority of the studies the rate was <10%. In two trials the rate was higher (range 19–42%).

#### Venous thromboembolism

The numbers of randomly assigned and analyzed patients in each included trial are described in Table I. Seven trials reported on the primary outcome of our meta-analysis, symptomatic VTE. LMWH significantly reduced symptomatic VTE [RR 0.46 (95% CI  $(0.32-0.67, I^2-0.6\%)$ ] as presented in Figure 2a. Ten trials (6942 patients) were eligible for meta-analysis of any VTE [17-19,21-27]. LMWH significantly reduced any VTE [RR 0.56 (95% CI 0.38–0.81,  $I^2$ -36%]. We specifically appraised the risk for DVT and PE. LMWH significantly reduced symptomatic DVT [RR 0.35 (95% CI 0.21-0.61)] and PE [RR 0.49 (95% CI 0.29-0.84)] as presented in Figure 2b and c. In a subgroup analysis of VTE in patients with lung cancer LMWH further reduced VTE [RR 0.42] (95% CI 0.25–0.71)], as depicted in Figure 3a. The striking effect of LMWH on VTE reduction has been documented in the subgroup of patients with pancreatic cancer [RR 0.31 (95% CI 0.18-0.55)]. VTE analyses are depicted in Figure 3b. Sensitivity analysis according to risk of bias, and specifically according to allocation concealment showed similar results for VTE reduction in both trials of low risk for bias (RR 0.67, 95% CI 0.49–0.93) and those of high risk (RR 0.33, 95% CI 0.21–0.52).

#### All-cause mortality

Six trials (2550 patients) reported on all-cause mortality. All trials reported mortality rates at 12 months of follow-up; five trials reported mortality at six months and four trials reported it also at 24 months. LMWH had no significant effect on survival (at 12 months [RR 0.93 (95% CI 0.83–1.04)]. RR for mortality at other time points resembled this figure. Survival RRs are depicted in Figure 4. We lacked the mortality data for subgroup of lung cancer patients and pancreatic cancer patients.

First Author, Year [Ref]	Design	LMWH, Schedule	Duration of Treatment	Number of Patients	Cancer type	Disease Stage	Concomitant therapy
Altinbas, 2004 [17] Kakkar 2004 [18]	RCT Open label Prospective, multicenter RCT	(5000 IU sc, od) Control = no Dalteprin Dalteparin (5000 IU sc, od) Control = placebo	18 weeks 1 year	42 LMWH 42 Control 196 LMWH 189 Control	Small cell lung carcinoma Breast, lung, gastrointestinal, pancreas, liver, genitourinary, ovary, or uterus	Metastatic or LA Metastatic or LA	Chemotherapy Chemotherapy and/or radiotherapy
Klerk 2005 [19]	RCT	Nadroparin received body weight-adjusted therapeutic doses of subcutaneous nadroparin for 2 weeks (<50 kg, 3,800 IU twice daily; 50–70 kg, 11 400 IU once daily;>70 kg, 15 200 IU once daily) followed by half-therapeutic doses for an additional 4 weeks (<50 kg, 3,800 IU once daily;>70 kg, 5,700 IU once daily;>70 kg, 7,600 IU once daily)	6 weeks	148 LMWH 154 Control	Solid cancers	Metastatic or LA	Chemotherapy and/or radiotherapy
Sideras 2006 [20]	RCT	Dalteparin (5000 IU sc, od) Control = placebo	18 weeks or up to disease progression	71 LMWH 70 Control	Breast cancer, prostate cancer, lung cancer, colorectal cancer	Metastatic	Chemotherapy and/or radiotherapy
Agnelli 2009 [PROTECHT; 21]	Prospective, multicenter RCT	Nadroparin (3800 IU sc, od) Control = placebo	Duration of chemotherapy or up to a maximum of 120 days	779 LMWH 387 Control	Gastrointestinal, pancreatic, breast, ovarian, or head and neck cancer	Metastatic or LA	Chemotherapy
Riess 2010 [CONKO004; 22]	RCT Open label	Enoxaparin (1mg/kg, sc, od) Control = no Enoxaparin	12 weeks	160 LMWH 152 Control	Pancreatic cancer	Metastatic or LA	Chemotherapy
Van Doormaal 2011 [23]	RCT	Nadroparin: received body weight-adjusted subcutaneous therapeutic doses for 2 weeks (< 50 kg, 3,800 IU twice daily; 50-70 kg, 11 400 IU once daily;>70 kg, 15 200 IU once daily;>70 kg, 15 200 IU once daily;>70 kg, 15 200 iU once daily;>70 kg, 00 IU once daily; 50-70 kg, 5,700 IU once daily;>70 kg, 7,600 IU once daily)	12 weeks	244 LMWH 259 Control	Prostate cancer, lung cancer, pancreatic cancer	Metastatic or LA	Chemotherapy
Agnelli 2012 [SAVE ONCO; 24]	RCT	Semuloparin, 20 mg, sc, od)	3.5 months (median) Until change of chemotherapy	1608 LMWH 1604 Control	Lung cancer, pancreatic cancer, gastric cancer, colorectal cancer, bladder cancer, ovarian cancer	Metastatic or LA	Chemotherapy
Maraveyas 2012 [25]	RCT	Dalteparin 200 IU/kg sc, od for 4 weeks followed by a step down to 150 IU/kg for a further 8 weeks) Control = no Dalteparin	12 week	63 LMWH 60 Control	Pancreatic cancer	Metastatic or LA	Chemotherapy
Haas 2012 [TOPIC 1; 26,27]	RCT	Certoparin (3000 IU sc, od) Control = placebo	6 months	174 LMWH 177 Control	Breast cancer	LA	Chemotherapy
Hass 2012 [TOPIC 2; 26,27]	RCT	Certoparin (3000 IU sc, od) Control = placebo	6 months	273 LMWH 274 Control	Lung cancer	Metastatic or LA	Chemotherapy

Table I. Characteristics of studies included in the meta-analysis.

Abbreviations: IU, international units; SC, subcutaneous; LA, locally advanced; OD, once daily; RCT, randomized controlled trials.

#### Type of LMWH

There were no significant variations in the effect of the different members of the LMWH utilized in the included trials in any of the outcomes.

#### Adverse events

The rate of grade 3 or 4 adverse events was reported in nine trials, evaluating 6595 patients. There was no significant increase in either the rate of clinically relevant bleeding [RR 1.29 (95% CI 0.95–1.77)],

(A)	LM	NH	Co	ntrol		Risk ratio	Risk ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.6.2 semuloparin							
Agnelli 2012 (SAVE ONC Subtotal (95% CI)	O) 20	1608 1608	55	1604 1604	51.1% 51.1%	0.36 [0.22, 0.60] 0.36 [0.22, 0.60]	•
Total events	20		55				
Heterogeneity: Not applic Test for overall effect: Z =	able : 3.92 (F	P < 0.00	01)				
1.6.3 certoparin							
Haas 2012 (TOPIC 1)	3	174	4	178	6.0%	0.77 [0.17, 3.38]	
Haas 2012 (TOPIC 2) Subtotal (95% CI)	5	268 442	10	264 442	11.7% 17.7%	0.49 [0.17, 1.42] 0.57 [0.24, 1.36]	•
Total events	8		14				
Heterogeneity: 1 au <sup>2</sup> = 0.0 Test for overall effect: Z = 1.6.4 dalteparin	i0; χ² = i : 1.27 (p	0.23, df = 0.20)	= 1 (p = )	= 0.63)	; 1² = 0%		
Altinbas 2004	1	42	0	42	1.3%	3 00 [0 13 71 61]	
Kakkar 2004 (FAMOUS)	4	190	5	184	7.8%	0.77 [0.21, 2.84]	
Marvayas 2012	5	59	13	62	14.0%	0.40 [0.15, 1.06]	
Sideras 2006 Subtotal (95% CI)	4	68 359	5	69 357	8.1% 31.2%	0.81 [0.23, 2.89] 0.62 [0.32, 1.19]	•
Total events	14		23				
Heterogeneity: Tau <sup>2</sup> = 0.0	10; χ <sup>2</sup> =	1.99, df	= 3 (p =	= 0.58)	; I <sup>2</sup> = 0%		
Test for overall effect: Z =	1.45 (p	= 0.15)	)				
Total (95% CI)		2409		2403	100.0%	0.46 [0.32, 0.67]	•
Total events	42		92				
Heterogeneity: Tau <sup>2</sup> = 0.0 Test for overall effect: Z =	10; χ <sup>2</sup> = 4 4.14 (p	4.11, df	= 6 (p = 01)	= 0.66)	; I <sup>2</sup> = 0%	0.01	0.1 1 10 1
Test for subgroup differer	10065' V <sup>2</sup>	= 1 90	df = 2 (	n = 0.3	(9) I <sup>2</sup> = 0	%	ours Livivvmravours cont

nor in major bleeding events [RR 1.28 (95% CI 0.84–1.95)]. There was no significant increase in thrombocytopenia [RR 1.05 (95% CI 0.76–1.45)]. We could not conduct subgroup analysis for bleeding according to type of malignancy since data were not provided separately.

#### Discussion

The results of our meta-analysis indicate that administration of LMWH as primary thrombopro-

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	LMV	νH	Cor	ntrol		Risk	ratio	Ris	k ratio	
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Rand	dom, 95% CI	M-H, Ran	dom, 95%	CI
1.9.2 semuloparin										
Agnelli 2012 (SAVE ONC Subtotal (95% CI)	0) 11	1608 1608	34	1604 1604	62.7% 62.7%	0.32 0.32	[0.16, 0.63] [0.16, 0.63]	-		
Total events	11		34							
Heterogeneity: Not applic	able									
Test for overall effect: Z =	3.28 (p	= 0.00	1)							
1.9.3 certoparin										
Haas 2012 (TOPIC 1)	2	174	4	178	10.1%	0.51	[0.09, 2.76]			
Haas 2012 (TOPIC 2)	4	268	9	264	21.1%	0.44	[0.14, 1.40]		-	
Subtotal (95% CI)		442		442	31.2%	0.46	[0.18, 1.20]	-		
Total events	6		13							
Heterogeneity: Tau <sup>2</sup> = 0.0	0; χ <sup>2</sup> = 0	.02, df	= 1 (p =	0.88);	I <sup>2</sup> = 0%					
Test for overall effect: Z =	1.59 (p	= 0.11	)							
1.9.4 dalteparin										
Kakkar 2004 (FAMOUS)	1	190	4	184	6.0%	0.24	[0.03, 2.15]		_	
Subtotal (95% CI)		190		184	6.0%	0.24	[0.03, 2.15]			
Total events	1		4							
Heterogeneity: Not applic	able									
Test for overall effect: Z =	1.27 (p	= 0.20	)							
Total (95% CI)		2240		2230	100.0%	0.35	[0.21, 0.61]	•		
Total events	18		51							
Heterogeneity: Tau <sup>2</sup> = 0.0	$0: \gamma^2 = 0$	.50. df	= 3 (p =	0.92)	l <sup>2</sup> = 0%					
Test for overall effect: Z =	3.80 (p	= 0.00	01)	. ,			0.01	0.1 1	I 10	100
Test for subgroup differen	nces: χ <sup>2</sup> :	= 0.48,	df = 2 (p	o = 0.7	9), I² = 0	%	Fav		nravours o	Untrol

(C) Study or Subgroup Events Total Events Total Weight M-H, Random, 95% Cl M-H, Random, 95% Cl

1.8.1 nadroparin									
Agnelli 2009 (PROTECHT)	3	769	3	381	11.6%	0.50 [0.10, 2.44]		-	
van Doormaal 2011	3	244	7	259	16.4%	0.45 [0.12, 1.74]			
Subtotal (95% CI)		1013		640	28.0%	0.47 [0.17, 1.32]	-		
Total events	6		10						
Heterogeneity: Tau <sup>2</sup> = 0.00; γ	ζ² = 0	0.01, df =	: 1 (p =	= 0.94)	I <sup>2</sup> = 0%				
Test for overall effect: Z = 1.4	14 (p	= 0.15)							
1.8.2 semuloparin									
Agnelli 2012 (SAVE ONCO)	10	1608	24	1604	54.6%	0.42 [0.20, 0.87]			
Subtotal (95% CI)		1608		1604	54.6%	0.42 [0.20, 0.87]	-		
Total events	10		24						
Heterogeneity: Not applicable	Э								
Test for overall effect: Z = 2.3	34 (p	= 0.02)							
1.8.3 certoparin									
Haas 2012 (TOPIC 1)	1	174	1	178	3.9%	1.02 [0.06, 16.23]			
Haas 2012 (TOPIC 2)	2	268	4	264	10.3%	0.49 [0.09, 2.67]		-	
Subtotal (95% CI)		442		442	14.2%	0.60 [0.14, 2.54]	-		
Total events	3		5						
Heterogeneity: Tau <sup>2</sup> = 0.00; γ	ζ² = C	0.20, df =	: 1 (p =	= 0.66)	l <sup>2</sup> = 0%				
Test for overall effect: Z = 0.6	69 (p	= 0.49)							
1.8.4 dalteparin									
Kakkar 2004 (FAMOUS)	2	190	0	184	3.2%	4.84 [0.23, 100.20]	-+-	-	
Subtotal (95% CI)		190		184	3.2%	4.84 [0.23, 100.20]			
Total events	2		0						
Heterogeneity: Not applicable	Э								
Test for overall effect: Z = 1.0	02 (p	= 0.31)							
Total (95% CI)		3253		2870	100.0%	0.49 [0.29, 0.84]	•		
Total events	21		39						
Heterogeneity: Tau <sup>2</sup> = 0.00; y	(² = 2	.69, df =	5 (p =	= 0.75)	I <sup>2</sup> = 0%	0.01	01 1	10	100
Test for overall effect: Z = 2.5	57 (p	= 0.01)				Fav	ours LMWHF	avours d	control
Test for subgroup differences	s: χ² :	= 2.47, d	f = 3 (	p = 0.4	8), I <sup>2</sup> = 0%				

Figure 2. Forest plot of risk ratios (RRs) comparing (A) Symptomatic venous thromboembolism (VTE) (B) Deep vein thrombosis symptomatic (DVT) and (C) pulmonary embolism for patients who received LMWH in addition to standard therapy versus those who received standard therapy only. Risk ratios for each trial are represented by the **squares**, the size of the square represents the weight of the trial in the meta-analysis, and the **horizontal line** crossing the square represents the 95% confidence interval (CI). The **diamonds** represent the estimated overall effect based on the meta-analysis random effects of all trials.

(A)	LMWH	4	Conti	rol		Risk ratio	Risk ratio
Study or Subgroup	Events 1	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
1.10.1 nadroparin							
Agnelli 2009 (PROTECHT Subtotal (95% CI)	) 7	199 199	7	80 80	26.1% 26.1%	0.40 [0.15, 1.11] 0.40 [0.15, 1.11]	•
Total events	7		7				
Heterogeneity: Not applica	ible						
Test for overall effect: Z =	1.76 (p =	0.08)					
1.10.2 semuloparin							
Agnelli 2012 (SAVE ONCO Subtotal (95% CI)	)) 9	591 591	25	589 589	47.3% 47.3%	0.36 [0.17, 0.76] 0.36 [0.17, 0.76]	•
Total events	9		25				
Heterogeneity: Not applica	ble						
Test for overall effect: Z =	2.67 (p =	800.0	)				
1.10.3 certoparin							
Haas 2012 (TOPIC 2)	5	268	10	264	23.9%	0.49 [0.17, 1.42]	
Subtotal (95% CI)		268		264	23.9%	0.49 [0.17, 1.42]	-
Total events	5		10				
Heterogeneity: Not applica	.ble						
Test for overall effect: Z =	1.31 (p =	0.19)					
1.10.4 dalteparin							
Altinbas 2004	1	42	0	42	2.7%	3.00 [0.13, 71.61]	
Subtotal (95% CI)		42		42	2.7%	3.00 [0.13, 71.61]	
Total events	1		0				
Test for succell offersts 7 -		0.50					
Test for overall effect: Z =	0.66 (p =	0.50)					
Total (95% CI)		1100		975	5 100.0%	0.42 [0.25, 0.71]	•
Total events	22		42				
Heterogeneity: Tau <sup>2</sup> = 0.00	); χ <sup>2</sup> = 1.7	4, df :	= 3 (p =	0.63);	I <sup>2</sup> = 0%	H-	
Test for overall effect: Z =	3.26 (p =	0.001	)			E:	avours I MWHEavours control
Test for subgroup difference	ces: χ² = 1	1.74, 0	df = 3 (p	= 0.63	8), I <sup>2</sup> = 0%		
(B)	LMWF	4	Conti	rol		Risk ratio	Risk ratio
Study or Subgroup	Events 7	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% Cl
1.14.1 nadroparin							
Agnelli 2009 (PROTECHT Subtotal (95% CI)	) 3	36 36	3	17 17	14.3% 14.3%	0.47 [0.11, 2.10] 0.47 [0.11, 2.10]	
Total events	3		3				
Heterogeneity: Not applica	ible						
Test for overall effect: Z =	0.99 (p =	0.32)					
1.14.2 semuloparin							
Agnelli 2012 (SAVE ONCO Subtotal (95% CI)	D) 3	126 126	14	128 128	21.4% 21.4%	0.22 [0.06, 0.74]	•
Total events	3		14				
Heterogeneity: Not applica	ble						
Test for overall effect: Z =	2.44 (p =	0.01)					
1.14.3 dalteparin							
Marvayas 2012	7	59	17	62	49.3%	0.43 [0.19, 0.97]	
Subtotal (95% CI)		59		62	49.3%	0.43 [0.19, 0.97]	-
Total events	7		17				

cancer LMWH reduced the RR for VTE even further [0.28 (95% CI 0.16-0.49) for pancreatic cancer] and [0.43 (95% CI 0.26-0.71) for lung cancer]. LMWH had no effect on survival in meta-analysis of all trials.

Based on our meta-analysis the NNT to prevent one symptomatic VTE is 50 (95% CI 33-100). Among lung cancer patients the NNT to prevent one VTE is 33 (95% CI 25-100) and in pancreatic cancer the NNT to prevent one VTE is 10 patients (95% CI 7-16), indicating a substantial benefit for LMWH in this subpopulation. The rate of serious adverse events was low, with the number needed to harm (NNH) being 100 (95% CI 50-very large number) for clinically relevant bleeding. The RR for major bleeding events was not greater compared with the control arm and neither the rate of thrombocytopenia. We could not infer based upon the analysis whether a specific LMWH exhibit superior results compared with other agents of this group, it is therefore probably safe to assume there is a class effect.

The rationale for primary thromboprophylaxis in cancer patients arises from the marked risk of cancer-associated VTE. Population-based case-control studies indicate a two-year cumulative incidence of 0.6-7.8%, depending on the population studied [28,29]. The risk for VTE depends profoundly on the primary site of cancer, whereas pancreatic, gastric and lung cancers confer the highest risk to develop VTE [30]. Lung and cardiac comorbidities which are frequent in lung cancer patients increase the risk of VTE by 20% [31,32]. Former studies indicate that the incidence of VTE is highest within the first six months of commencing the anti-cancer treatment [2].

Our meta-analysis did not show a survival advantage. Several studies have indicated that selected populations may gain a survival advantage from LMWH prophylaxis, whereas the LMWH benefit was most apparent among patients with a better prognosis. Some other considerations that may impact survival analysis include the short length of follow-up: in some of the studies in a subgroup analysis, the good-prognosis group of patients experienced a superior survival with LMWH [18,19]. Another determinant is the effect of LMWH in different tumor types and disease stages. The majority of studies included in the meta-analysis encompass a variety of tumor types whereas the biological role of LMWH may differ in distinctive cancers. Due to the main role of coagulation pathways in pancreatic and lung cancer, the potential benefit of concomitant administration of LMWH may be enhanced in these cancers, as we have shown in this meta-analvsis in terms of VTE reduction in these entities.

Several limitations of this analysis must be acknowledged. The heterogeneity of cancer types and disease stages in some of the studies may attenuate the impact

( )							
Total events	3		14				
Heterogeneity: Not applicable							
Test for overall effect: Z = 2.4	4 (p =	0.01)					
1.14.3 dalteparin						_	
Marvayas 2012	7	59	17	62	49.3%	0.43 [0.19, 0.97]	
Subtotal (95% CI)		59		62	49.3%	0.43 [0.19, 0.97]	
Total events	7		17				
Heterogeneity: Not applicable							
Test for overall effect: Z = 2.0	4 (p =	0.04)					
1.14.4 enoxaparin							
Riess 2010 (CONKO004)	2	160	15	152	15.0%	0.13 [0.03, 0.54]	
Subtotal (95% CI)		160		152	15.0%	0.13 [0.03, 0.54]	
Total events	2		15				
Heterogeneity: Not applicable							
Test for overall effect: Z = 2.7	8 (p =	0.005)					
						•	
Total (95% CI)		381		359	100.0%	0.31 [0.18, 0.55]	
Total events	15		49				
Heterogeneity: Tau <sup>2</sup> = 0.00; $\chi$	² = 2.8	i8, df = 3	(p = 0	.41); l <sup>a</sup>	<sup>2</sup> = 0%		-
Test for overall effect: Z = 4.0	1 (p <	0.0001)				Eavours I MWHE	a
Test for subgroup differences	$\gamma^2 = 2$	2.73 df :	= 3 (n =	= 0.441	$1^2 = 0\%$	T drouio Emitin	-

Figure 3. Forest plot of risk ratios (RRs) comparing (A) venous thromboembolism (VTE) in lung cancer patients and (B) in pancreatic cancer patients who received LMWH in addition to standard therapy versus those who received standard therapy only. Risk ratios for each trial are represented by the squares, the size of the square represents the weight of the trial in the meta-analysis, and the horizontal line crossing the square represents the 95% confidence interval (CI). The diamonds represent the estimated overall effect based on the meta-analysis random effects of all trials.

10 100

phylaxis to cancer patients concomitantly with standard chemotherapy significantly reduces the risk for symptomatic VTE, any VTE and PE, while the risk for major bleeding is not significantly increased. In a subgroup analysis of pancreatic cancer and lung

(A)	LM	WH s Tota	Cont Events	trol	Weight	Risk ratio	Risk ratio
1.1.1nadroparin	Litent	0 1000	Litente	1010	in ricigin	m-ri, rixed, march	
Klerk 2005 Subtotal (95% Cl)	5	8 148 14	52	15	4 6.0%	1 16 [0 86, 1 56] 1 16 [0 86, 1 56]	ţ.
Total events Heterogeneity Not applicat Test for overall effect. Z = 0	5 ole 98 (p = 0.3	8	52				
112semuloparin							
Agnelli 2012 (SAVE ONCO) Subtotal (95% CI)	69	8 1608 160	8 714	160	4 84.7%	0.98 [0.90, 1.05]	
Total events Heterogeneity Not applicat Test for overall effect: Z = 0	69 63 (p = 0.5	8 53)	714				
1.1.3 certoparin							
Haas 2012 (TOPIC 1) Haas 2012 (TOPIC 2) Subtotal (65% Ch	1 5	5 174	12	17	8 1.496 3 7.096	1 28 [0.62, 2.65] 0.93 [0.67, 1.29]	Ŧ
Total events Heterogeneity $\chi^2 = 0.60$ , df Test for overall effect: Z = 0	7 = 1 (p = 0 06 (p = 0.5	0 44),   <sup>2</sup> =	71 096			0.57[0.13,2.23]	Ĭ
1.1.4 dalteparin							
Marvayas 2012 Subtotal (95% CI)	3	4 54	7	6	2 0.8%	0.60 [0.19, 1.95] 0.60 [0.19, 1.95]	-
Total events Heterogeneity: Not applical Test for overall effect: Z = 0	ble 85 (p = 0.4	4	7				
Total (95% Ch Total events	83	226	844	227	1 100.04	0.98[0.91,1.06]	
Test for overall effect: Z = 0 Test for subgroup difference	41 (p = 0.6 es χ <sup>2</sup> = 1	.90, df =	3 (p = 0.59	7), 1² = 0	990		0.01 0.1 1 10 100 Favours LMWH Favours control
(B)	I MW		Contro	4		Risk ratio	Risk ratio
Study or Subgroup	Events	Total	Events	Total	Weight	M-H, Random, 95% CI	M-H, Random, 95% CI
Accelli 2008 (PROTECHT)	11*	740		201	22.455	1 04 10 02 1 31	
Klerk 2005	333	148	112	154	19 796	0.84 (0.71, 0.98)	-
van Doormaal 2011 Subtotal (95% Ch	138	244	160	259	21.4% 62.6%	0.92 (0.79, 1.06) 0.94 (0.81, 1.08)	1
Van Doormaal 2011 Subtotal (95% CI)	90 138	148 244 1161	112	154 259 794	19.7% 21.4% 62.6%	0.84 [0.71, 0.98] 0.92 [0.79, 1.08] 0.94 [0.81, 1.08]	

Subtotal (95% CI)		1161		794	62.6%	0.94 (0.81, 1.08)		•		
Total events	561		427							
Heterogeneity Tau <sup>2</sup> = 0.01, )	(2 = 5.20, dt	f = 2 (p =	0.07), 1 <sup>2</sup>	= 6296						
Test for overall effect: Z = 0.9	1 (p = 0.36)									
1.2.+dalteparin										
Altinbas 2004	20	42	30	42	7.090	0.67 [0.46, 0.97]		-		
Kakkar 2004 (FAMOUS)	105	190	112	184	18.690	0.91 [0.76, 1.08]				
Sideras 2006	45	68	41	69	11.895	1 11 [0 86, 1.44]		1		
Subtotal (95% CI)		300		295	37.4%	0.90[0.71,1.14]		•		
Total events	170		183							
Heterogeneity Tau <sup>2</sup> = 0.03, 3	(= 5.04, dt	f=2(p=	0.08), 12	= 6090						
Test for overall effect. Z = 0.8	4 (p = 0.40)									
Total (95% Ch		1461		1089	100.0%	0.93[0.83,1.04]		•		
Total events	731		610							
Heterogeneity Tau <sup>2</sup> = 0.01, 3	(" = 10.19, 0	df = 5 (p	= 0.07), 1	= 5190				_	<u>+</u>	
Test for overall effect. Z = 1.3	3 (p = 0.18)						Eavours I	MMH Eav	10	trol
Test for subgroup difference	s X2 = 0.00	s, df = 1 (	p = 0.80)	12 = 04			1 avours L	annin rar	ours con	

Figure 4. Forest plot of risk ratios (RRs) comparing (A) mortality at 6 months or (B) 12 months for patients who received LMWH in addition to standard therapy versus those who received standard therapy only. Risk ratios for each trial are represented by the **squares**, the size of the square represents the weight of the trial in the meta-analysis, and the **horizontal line** crossing the square represents the 95% confidence interval (CI). The **diamonds** represent the estimated overall effect based on the meta-analysis random effect of all trials.

of LMWH, especially on survival outcomes. We could only conduct subgroup analysis for VTE reduction for lung and pancreatic cancers, but not other types of malignancies. Moreover, there were various LMWHs used and in different dosages. The short follow-up is a flaw of the included studies. As mentioned, post hoc analysis of longer follow-up in some of the studies revealed a survival benefit in specific populations, yet meta-analysis of these figures was not feasible.

Future research should focus on identifying risk factors that predispose patients to develop VTE in order to define the subgroup of patients that would benefit the most from the addition of LMWH to the treatment regimen, considering cancer type and stage. The type of LMWH preparation and the optimal prophylactic dosage should be further explored. Future study design should be powered to evaluate a survival endpoint, but also quality of life parameters in patients for whom expected survival is poor. In addition, nowadays new oral anticoagulants (NOAC) are available. They are both convenient due to the oral administration and safe. These agents have been extensively studied for prophylaxis of acute VTE, long-term anticoagulation for atrial fibrillation, and acute coronary syndromes [33]. Future trials should assess their role in prophylaxis in ambulatory cancer patients.

Cost effectiveness analysis should be appraised as well, since most of the included studies did not report hospitalization rate and cost. It is important to investigate whether specific LMWHs are superior to other agents since there are data suggesting that some of the antitumor effects of LMWH are dependent on molecular weight.

In conclusion, our meta-analysis shows a significant reduction in all types of VTE, with no apparent increase in the incidence of major bleeding episodes. The reduction in VTE may translate into improved quality of life, less hospitalizations and consequently fewer delays in the administration of chemotherapy. Our analysis indicates an exceptionally marked benefit for patients with pancreatic cancer where NNT for VTE to prevent one VTE is 10 patients (95% CI 7–16) and in lung cancer patients [NNT 33 (95% CI 25–100)]. Therefore, our data suggest that practitioners should strongly consider the use of LMWH as primary thromboprophylaxis to significantly reduce the rate of VTE specifically in pancreatic cancer patients, and most possibly lung cancer patients.

**Declaration of interest:** The authors report no conflicts of interest. The authors alone are responsible for the content and writing of the paper.

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#### Supplementary material available online

Supplementary Table I. (available online at http:// informahealthcare.com/doi/abs/10.3109/0284186X. 2014.934397). boembolic events in ambulatory patients with metastatic or locally advanced solid cancer receiving chemotherapy: A randomised, placebo-controlled, double blind study. Lancet Oncol 2009;10:943–9.

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